

Andrew J. Spieker, PhD
BIOS 6312 - Modern Regression Analysis
Spring 2020
Exam #1 Key

Instructions: Please adhere to the following guidelines:

- This exam is an individual effort (closed book/notes/phone/everything), although you are permitted the use of a scientific calculator.
 - Please read the questions carefully and answer only what you are being asked to answer.
 - Please be concise. You are asked many times throughout the exam to state something in no more than a sentence or two, meaning you should be able to provide the answer in the amount of space provided.
 - See me if you need scratch paper.
 - Do not spend too much time on one problem. Mind the point distribution (below).
 - Please write legibly.
 - Please round any final calculations to **three** significant digits!
 - There are ten numbered pages (seven required problems plus two optional problems). There are five pages of appendix material. Check to make sure you have everything before beginning.
-

Score Range	Frequency
95+	2
90-94	5
85-89	4
80-84	4
75-79	1
70-74	2

Statistics	
Mean	85.8%
Median	85.5%
Corr($\overline{HW}_{1:3}$, Exam 1)	0.603

These scores *do not* reflect the optional points, which will be accounted for at the end of the semester.

1. A group of co-investigators is using data from a cross-sectional study ($N = 46$) to obtain insights into the association between age and the prostate-specific antigen (PSA, in ng/ml) among an adult male population. Two of the investigators on the team, Michael and Rebecca, take it upon themselves to each perform an analysis of these data using simple linear regression in Stata (Appendix I). When presenting results at the team meeting, they find that their approaches differ in that Rebecca log-transformed PSA and altered the units of age from “years” to “decades” by dividing the `age` variable by ten.

(a) Circle the letter corresponding to the statement with which you most closely agree:

- A. We should use Rebecca’s analysis if we have evidence that PSA is right-skewed.
- B. Since linear regression requires normal errors, we should look at a quantile-quantile plot of the studentized residuals from Rebecca’s and Michael’s analyses before deciding which to use.
- C. Robust standard errors serve no purpose in Rebecca’s analysis because log-transformation always eliminates heteroscedasticity.
- D. A good reason to use the results of Michael’s analysis would be if the investigators pre-specified that they would examine the association between age and *mean* PSA.
- E. Since Rebecca’s analysis does *not* achieve statistical significance (see Appendix I), we should instead publish Michael’s analysis and forget about Rebecca’s altogether.

(b) Use the Stata output of one of the analyses (provided in Appendix I) to briefly summarize this study’s evidence regarding an association between age and *geometric* mean PSA. In your response, you need only include (with proper interpretation) a point estimate and its 95% confidence interval, as well as a summary of your conclusion with appropriate inferential measures.

Ans: These data do not provide sufficient evidence of such an association ($p = 0.057$). We estimate that, when comparing subgroups differing in their age by ten years, the older subgroup has a 20.4% higher geometric mean. Based on a 95% confidence interval, these data would not be judged unusual if in truth the older subgroup had anywhere between a 0.536% lower and 45.8% higher geometric mean. Alternatively, you could have expressed this as a geometric mean ratio and obtained a point estimate of 1.20 (95% CI: [0.995, 1.46]).

(c) Devon uses the regression output from Michael’s analysis to form a 95% prediction interval for PSA among 30 year-old men. After examining the (studentized) residual-versus-predictor plot provided in Appendix II, state the *two* most obvious reasons why this is an extraordinarily bad idea on Devon’s part. Do *not* attempt to form the prediction interval.

Ans: The first big issue is one of extrapolation; an age of 30 years is clearly outside the age range of these data (about 50 to 80 years, it seems). The second is that the residual-versus-predictor plot is suggests a violation to homoscedasticity, an assumption we would need in order to form a valid prediction interval based on the methods we have learned.

2. A cross-sectional observational study was conducted to determine the association between number of hours of sleep the night before an exam (X , in hours) and the subsequent exam score (Y , %). $N = 500$ students were randomly sampled from Vanderbilt University and were surveyed regarding their sleep history and most recent exam score. Once the data were collected, the following simple linear regression model was fit using ordinary least squares:

$$\mathbf{E}[Y|X = x] = \beta_0 + \beta_1 x.$$

Your response to sub-questions (a) through (d) below should each be **at most two sentences**.

- (a) Based on the scatterplot provided in Appendix III, are you concerned about any violations to the assumption of linearity for the purposes of estimating β_1 ?

Ans: The scatterplot does not seem to suggest any egregious violations to the assumption of linearity, and so I am not worried. Even if there were a departure from linearity, we have learned that the regression coefficient is still interpretable as a weighted average of all the pairwise slopes (i.e., we would be estimating some first-order trend).

- (b) Briefly describe how the scatterplot provided in Appendix III demonstrates graphical evidence of non-normality of the errors.

Ans: There is a clear “ceiling” (the maximum exam score is 100%), which we would not be able to see if the errors were normally distributed. Moreover, there is evidence of some left-skewness, particularly for lower values of sleep. Either of these examples is acceptable; you were only required to provide one.

- (c) You find out that one-hundred students in your sample were roommates with another student in your sample. Thinking in terms of whether the regression assumptions are met in this setting, state how this limits your ability to conduct inference on β_1 when using ordinary least squares.

Ans: We would be concerned, primarily because this would violate the assumption of independent errors (one person’s exam score may be correlated with his or her roommate’s even after having conditioned on sleep).

- (d) Suppose you had instead fit a no-intercept model, $\mathbf{E}[Y|X = x] = \beta x$. State the additional assumption that would be made by this restriction and briefly comment on whether it is reasonable.

Ans: This imposes the additional restriction that the mean exam score is zero among those not getting *any* sleep (and assuming that zero is the minimum score, this in essence means that lack of sleep *guarantees* a score of zero). This does not seem to be a reasonable assumption.

3. We have considered the REACH study for numerous examples in our notes and on homework problems. The REACH study was a randomized controlled trial that sought to evaluate the effect of a text-message intervention on hemoglobin A1c. A subset of the variables from this study include:

$$X = \begin{cases} 0 & \text{if assigned to control group} \\ 1 & \text{if assigned to REACH group} \end{cases},$$

$$Y = \text{A1c at twelve months (\%)}, \text{ and}$$

$$Z = \text{A1c at baseline (\%)}.$$

Consider the following two linear regression models:

$$\begin{aligned} \mathbf{E}[Y|X = x] &= \beta_0^* + \beta_1^*x, \text{ and} \\ \mathbf{E}[Y|X = X, Z = z] &= \beta_0 + \beta_1x + \beta_2z \end{aligned}$$

For this problem, you may assume that the trial is well conducted so that there is *no* systematic confounding, that each model is correctly specified, and that Z and Y are linearly correlated.

- (a) Which of the following four statements is true?

- I. $\beta_1^* = \beta_1$
- II. $\beta_1^* > \beta_1$
- III. $\beta_1^* < \beta_1$
- IV. It is impossible to tell without more information.

Answer: **I**

- (b) You employ ordinary least squares (OLS) linear regression to estimate each model. Which of the following four statements is true?

- I. $\widehat{\beta}_1^* = \widehat{\beta}_1$
- II. $\widehat{\beta}_1^* > \widehat{\beta}_1$
- III. $\widehat{\beta}_1^* < \widehat{\beta}_1$
- IV. It is impossible to tell without more information.

Answer: **IV**

- (c) Again based on ordinary least squares estimation, which of the following four statements is true?

- I. $\text{Var}(\widehat{\beta}_1^*) = \text{Var}(\widehat{\beta}_1)$
- II. $\text{Var}(\widehat{\beta}_1^*) > \text{Var}(\widehat{\beta}_1)$
- III. $\text{Var}(\widehat{\beta}_1^*) < \text{Var}(\widehat{\beta}_1)$
- IV. It is impossible to tell without more information.

Answer: **II**

- (d) In at most two sentences, state the purpose of adjusting for baseline A1c (i.e., the second model).

Ans: Baseline A1c would serve as a precision variable in this setting, reducing the variability associated with estimation of the treatment effect.

4. We have considered a study of forced expiratory volume (FEV) in children for numerous examples in our notes and on homework problems. The FEV study was a cross-sectional observational study that sought to evaluate whether smoking was associated with FEV, a measure of lung capacity in which higher values signify better lung function. A subset of the variables from this study include:

$$X = \begin{cases} 0 & \text{if a non-smoker} \\ 1 & \text{if a smoker} \end{cases},$$

$$Y = \text{FEV (L/sec), and}$$

$$Z = \text{Age (years).}$$

Consider the following two linear regression models:

$$\begin{aligned} \mathbf{E}[Y|X = x] &= \beta_0^* + \beta_1^*x, \text{ and} \\ \mathbf{E}[Y|X = X, Z = z] &= \beta_0 + \beta_1x + \beta_2z. \end{aligned}$$

You may assume for this problem that both models are correctly specified.

- (a) Which of the following is true regarding ordinary least squares estimation of β_1^* (i.e., from the unadjusted model), assuming homoscedasticity?
- I. It is exactly the same as a t -test assuming equal variances between smoking groups.
 - II. It is exactly the same as a t -test allowing unequal variances between smoking groups.
 - III. It is approximately the same as a t -test assuming equal variances between smoking groups.
 - IV. It is approximately the same as a t -test allowing unequal variances between smoking groups.

Answer: **I**

- (b) Which of the following is true regarding ordinary least squares estimation of β_1^* (i.e., from the unadjusted model), allowing heteroscedasticity?
- I. It is exactly the same as a t -test assuming equal variances between smoking groups.
 - II. It is exactly the same as a t -test allowing unequal variances between smoking groups.
 - III. It is approximately the same as a t -test assuming equal variances between smoking groups.
 - IV. It is approximately the same as a t -test allowing unequal variances between smoking groups.

Answer: **IV**

- (c) Again based on ordinary least squares estimation, which of the following four statements is true?
- I. $\text{Var}(\widehat{\beta}_1^*) = \text{Var}(\widehat{\beta}_1)$
 - II. $\text{Var}(\widehat{\beta}_1^*) > \text{Var}(\widehat{\beta}_1)$
 - III. $\text{Var}(\widehat{\beta}_1^*) < \text{Var}(\widehat{\beta}_1)$
 - IV. It is impossible to tell without more information.

Answer: **IV**

- (d) In at most two sentences, state the purpose of adjusting for age (i.e., the second model).

Ans: Age is likely a common cause of both smoking and FEV; adjusting for age helps control for this confounding.

5. This problem pertains to interpretation of regression parameters, particularly when dealing with common variable transformations. For each sub-question, there is a true-or-false question included pertaining to the model posed in that sub-question. Please circle the letter corresponding to your answer (no follow-up explanation is required). Then, you are asked to use plain language to interpret one of the parameters in the model, which means that you are *not* to include mathematical/statistical notation (e.g., $\mathbf{E}[Y|X = x]$); however, you are permitted to use notation to refer to variables (e.g., “ X ”), and you are permitted to use numbers/percentages (e.g., “10%”).

- (a) Consider the following regression model:

$$\mathbf{E}[Y|X = x] = \beta_0 + \beta_1 x.$$

T **F** When X increases by one unit, Y increases by β_1 units.

Provide a plain-language interpretation for $3\beta_1$.

Ans: This is the difference in mean Y comparing subgroups differing in their value of X by three units.

- (b) Consider the following regression model:

$$\mathbf{E}[\log(Y)|X = x] = \beta_0 + \beta_1 \log(x).$$

T **F** If Y denotes a measure of concentration, it absolutely *must* be log-transformed.

Provide a plain-language interpretation for $e^{\log(1.2)\beta_1}$.

Ans: This is the ratio of the geometric means of Y comparing subgroups differing in their value of X by 20%.

- (c) Consider the following regression model (assume X is categorical with levels 0, 1, and 2):

$$\mathbf{E}[Y|X = x] = \beta_0 + \beta_1 1(x = 1) + \beta_2 1(x = 2)$$

T **F** To determine if X and (mean) Y are associated, we could test $H_0 : \beta_1 = \beta_2 = 0$.

Provide a plain-language interpretation for β_1 .

Ans: This is the difference in mean Y comparing subgroups with $X = 1$ to those with $X = 0$.

(d) Consider the following regression model:

$$\mathbf{E}[Y|X = x] = \beta_0 + \beta_1(x - 10) + \beta_2(x - 10)^2$$

T F To determine if X and Y are associated, it is sufficient to test $H_0 : \beta_1 = 0$.

Provide a plain-language interpretation for β_1 .

Ans: This is the instantaneous rate of change of the conditional mean of Y (given X) at $X = 10$.

(e) Consider the following regression model:

$$\mathbf{E}[Y|X = x, Z = z] = \beta_0 + \beta_1x + \beta_2z + \beta_3xz.$$

T F β_0 denotes the mean of Y among the subgroup with $X = Z = 0$.

Provide a plain-language interpretation for β_3 .

Ans: This is the difference in the *association* between X and mean Y at fixed values of Z , comparing subgroups differing in their value of Z by one unit; more specifically, *association* refers to the difference in mean Y comparing subgroups differing in their value of X by one unit, but of the same value for Z . I would also accept if you simply said this quantifies the extent to which Z modifies the association between X and mean Y (or equivalently, the extent to which X modifies the association between Z and mean Y).

(f) Consider the following regression model:

$$\mathbf{E}[Y|X = x, Z = z, W = w] = \beta_0 + \beta_1x + \beta_2z + \beta_3w + \beta_4xz + \beta_5xw + \beta_6zw + \beta_7xzw$$

T F If the three-way interaction between X , Y , and Z is of most clinical interest, we need only estimate β from the reduced model $\mathbf{E}[Y|X = x, Z = z, W = w] = \beta xzw$.

Provide a plain-language interpretation for β_4 .

Ans: This quantifies the extent to which Z modifies the association between X and mean Y (or equivalently, the extent to which X modifies the association between Z and mean Y) among the subgroup for which $W = 0$.

6. A study was conducted to evaluate potential predictors of systolic blood pressure, (SBP, in mmHg, denoted by Y). As a simple example, consider alcohol consumption, X_1 (oz/week) and race category, X_2 , as predictors. The race variable in this example is coded as follows:

$$X_2 = \begin{cases} 1 & \text{White} \\ 2 & \text{Black} \\ 3 & \text{Asian} \\ 4 & \text{Other} \end{cases}$$

The following regression model is fit using Stata:

$$\mathbf{E}[Y|X_1 = x_1, X_2 = x_2] = \beta_0 + \beta_1 x_1 + \beta_2 1(x_2 = 2) + \beta_3 1(x_2 = 3) + \beta_4 1(x_2 = 4).$$

For this problem, refer to the Stata output in Appendix IV, which includes the regression output and an F -test of the hypothesis $H_0 : \beta_2 = \beta_3 = \beta_4 = 0$.

- (a) Report the mean squared error, with an interpretation that holds regardless of heteroscedasticity.

Ans: The mean squared error is $19.556^2 = 382 \text{ mmHg}^2$. This is an estimate of the average stratum-specific error variance (i.e., averaging over the joint density of the predictors: alcohol consumption and race).

- (b) In a sentence, either state your conclusions on whether there is evidence that mean SBP differs between black and Asian individuals (adjusted for alcohol consumption) **or** state that this cannot be done on the basis of the output and suggest a modification to the model that would permit this.

Ans: We are unable to do this based on the output provided; instead, we could accomplish this easily by choosing the reference category to be one of either “black” or “Asian.”

- (c) In a sentence, summarize your conclusions on whether there is evidence of an overall association between alcohol consumption and mean SBP, adjusted for race.

Ans: There is sufficient evidence of such an association ($p = 0.020$).

- (d) In a sentence, summarize your conclusions on whether there is evidence of an overall association between race and mean SBP, adjusted for alcohol consumption.

Ans: There is not sufficient evidence of such an association ($p = 0.169$).

7. This is a continuation of Problem 6. The investigators decided that they would like to allow for an interaction between alcohol consumption (X_1) and race category (X_2). For this problem, refer to the Stata output in Appendix V. It may help you to know that $Z_{0.975} = 1.96$.

- (a) Form a 95% prediction interval for SBP among white individuals with no history of alcohol consumption and briefly state the assumptions required for the prediction interval to be valid — you are *not* expected to evaluate the extent to which they hold.

Ans: First, note that $\hat{y} = \hat{\beta}_0$ in this problem. The prediction interval is hence given by $\hat{\beta}_0 + 1.96\hat{\sigma} = 129.4707 \pm 1.96 \times 19.584 = [91.1, 168]$ (in mmHg). We assume independent errors, linearity, and homoscedastic, normally distributed errors.

- (b) Examine the F -test results from the `testparm` command. In a sentence, and in plain language, what is the hypothesis being tested by this command?

Ans: This is testing whether there is a race-adjusted association between alcohol consumption and mean systolic blood pressure.

- (c) Note from the output of the `testparm` command that the critical value is $F(4, 727)$. Explain briefly where the numbers 4 and 727 come from.

Ans: The number 4 is the number of parameters being simultaneously tested; the number 727 is the sample size ($N = 735$) minus the number of parameters in the model ($K + 1 = 7 + 1 = 8$).

- (d) Glen states that because each of the the p-values corresponding to the three coefficients reported under `race#c.alcoh` is larger than 0.05, there is not sufficient evidence of an interaction between race and alcohol consumption. Is this argument valid? If not, explain in no more than two sentences what test you would prefer to conduct to evaluate evidence of such an interaction.

Ans: Glen's logic is flawed. Instead, we would (jointly) test to see if all three of the interaction terms in the model are simultaneously zero.

8. **Optional problem 1:** This is an optional problem — do not attempt it until you have completed the rest of the exam. A small bonus can be earned from a correct response.

Consider a randomized trial with two primary treatment groups, as denoted by X :

$$X = \begin{cases} 0 & \text{control} \\ 1 & \text{experimental treatment} \end{cases}$$

Further suppose that a subset of individuals receiving the experimental treatment (i.e., those with $X = 1$) also receive an *additional* treatment, denoted by Z :

$$Z = \begin{cases} 0 & \text{no additional treatment} \\ 1 & \text{additional treatment} \end{cases}$$

The investigators are interested in evaluating whether the supplemental treatment, Z , has an additional effect on mean Y beyond that of X alone. Without thinking it all the way through, a statistician on the study fits the following model:

$$\mathbf{E}[Y|X = x, Z = z] = \beta_0 + \beta_1x + \beta_2z + \beta_3xz.$$

In a couple of sentences, explain what will go wrong here and why. Then, briefly propose one or two alternative ways of answering the clinical question at hand. If you have time to propose two alternatives, briefly characterize the relative advantages and disadvantages of each of those two approaches.

Ans: Because $Z = XZ$ (collinearity, or more specifically in this case, *redundancy*), a Stata command in which you try to fit this model would not be able to provide output for all of the coefficients. Instead, the statistician could simply remove one of those two terms (Z , or XZ) and interpret the remaining coefficient as desired. Another option would be to fit the model: $\mathbf{E}[Y|Z = z] = \alpha_0 + \alpha_1z$ among the subset of people with $X = 1$. The advantage to the latter approach is that it does not make the assumption that X is associated with the (conditional) mean Y the same way in both the treatment and control groups ($X = 0$ and $X = 1$). If that *were* the case, the former approach could potentially provide improved efficiency.

9. **Optional problem 2:** This is an optional problem — do not attempt it until you have completed the rest of the exam. A small bonus can be earned from a correct response.

Let \mathbf{X} denote an $N \times (p + 1)$ design matrix of full rank. Recall that the hat matrix, \mathbf{H} , is defined as:

$$\mathbf{H} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T.$$

Prove that \mathbf{H} is an orthogonal projection matrix by first proving that it is symmetric (i.e., $\mathbf{H}^T = \mathbf{H}$) and then proving it is idempotent (i.e., that $\mathbf{H}^2 = \mathbf{H}$).

Since the ordinary least squares estimate is given by $\widehat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$, this in essence proves that $\widehat{\boldsymbol{\beta}}$ is estimated in a way such that $\widehat{\mathbf{y}} = \mathbf{X}\widehat{\boldsymbol{\beta}} = \mathbf{H}\mathbf{y}$ is the orthogonal projection of \mathbf{y} onto the linear subspace spanned by the columns of the design matrix \mathbf{X} .

Ans: First, $\mathbf{H} = (\mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T)^T = (\mathbf{X}^T)^T ((\mathbf{X}^T \mathbf{X})^{-1})^T (\mathbf{X})^T = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T = \mathbf{H}$. Moreover, we have $\mathbf{H}^2 = (\mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T)(\mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T) = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} (\mathbf{X}^T \mathbf{X})(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T = \mathbf{H}$.

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Appendix Material for Exam 1

APPENDIX I: Stata output for Problem 1

* MICHAEL'S ANALYSIS

```
. regress psa age, robust
```

```
Linear regression                Number of obs   =           46
                                F(1, 44)         =           7.93
                                Prob > F             =           0.0072
                                R-squared            =           0.1332
                                Root MSE         =           1.6947
```

```
-----+-----
              |               Robust
              |               Coef.   Std. Err.      t    P>|t|     [95% Conf. Interval]
-----+-----
              |
    age      |   .0821603   .029172     2.82   0.007   .0233679   .1409526
    _cons    |  -2.272699   1.795634   -1.27   0.212  -5.891561   1.346162
-----+-----
```

* REBECCA'S ANALYSIS

```
. gen logpsa = log(psa)
```

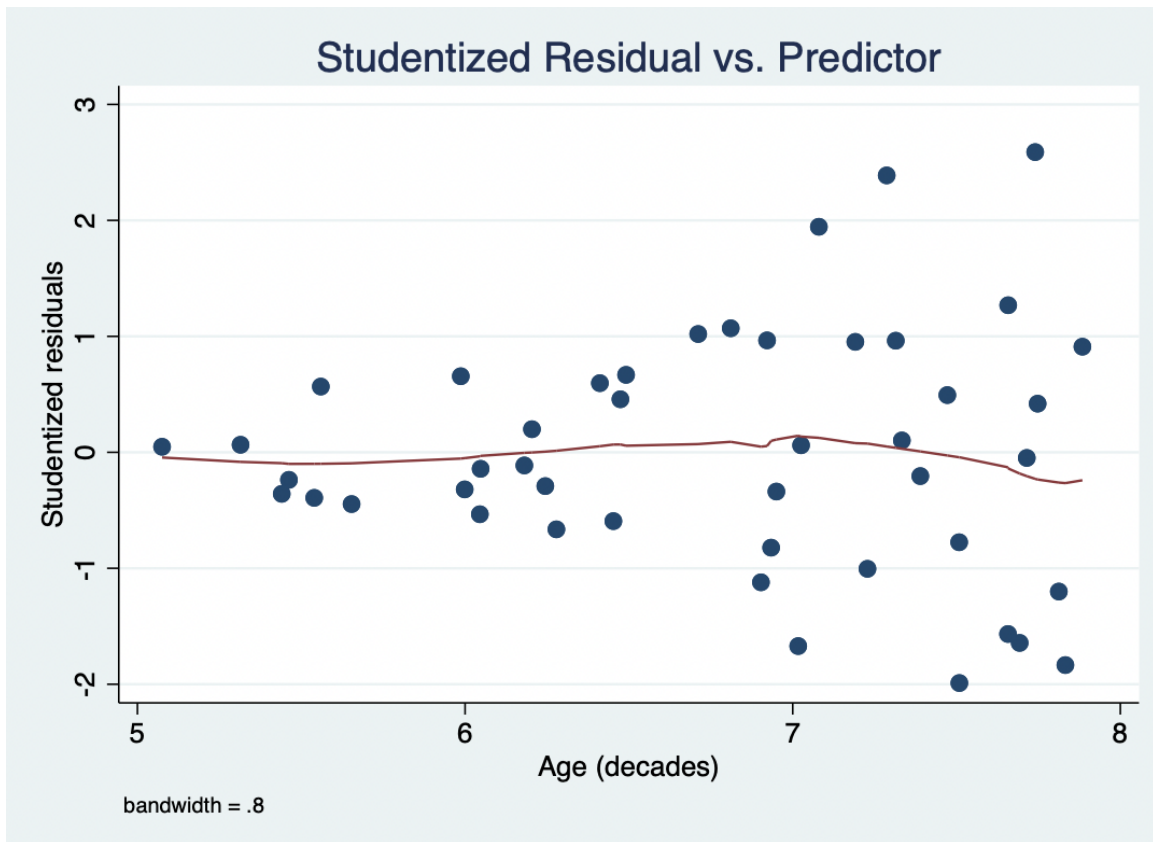
```
. gen age_decades = age/10
```

```
. regress logpsa age_decades, robust
```

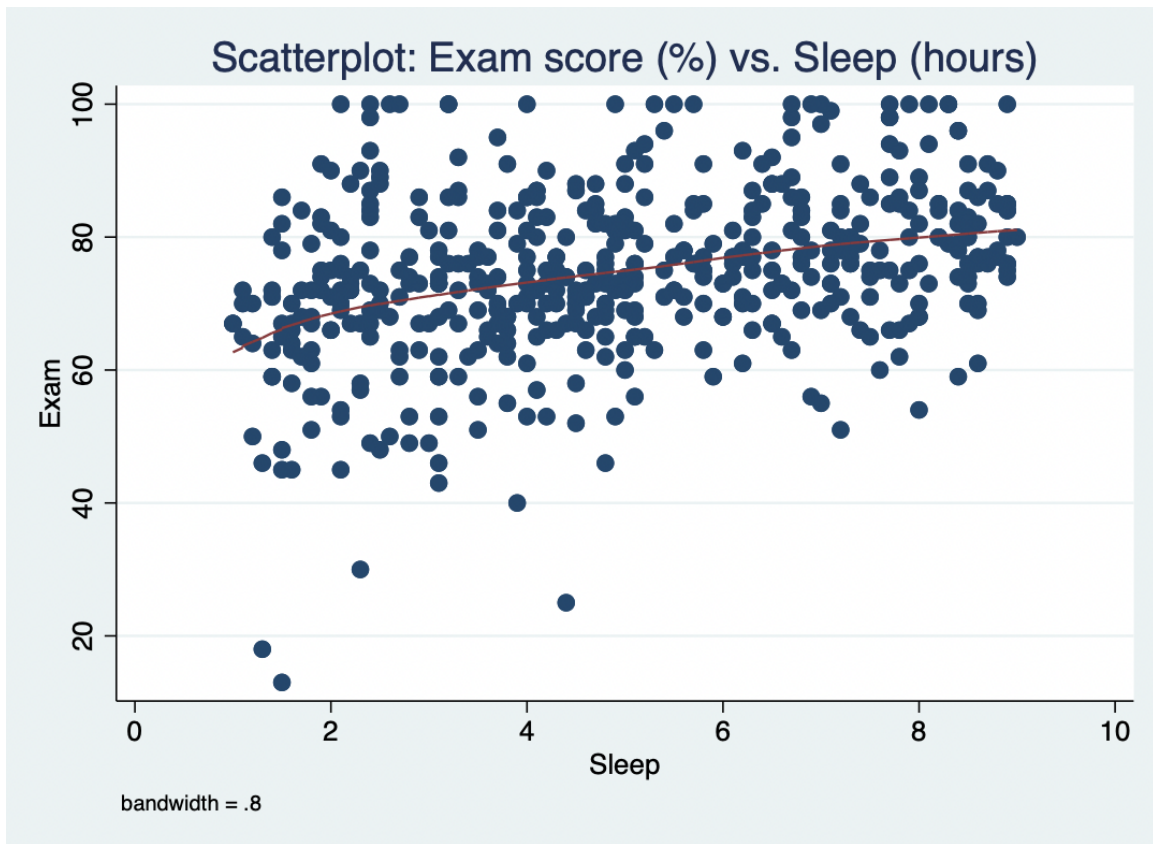
```
Linear regression                Number of obs   =           46
                                F(1, 44)         =           3.84
                                Prob > F             =           0.0565
                                R-squared            =           0.0662
                                Root MSE         =           .56421
```

```
-----+-----
              |               Robust
              |               Coef.   Std. Err.      t    P>|t|     [95% Conf. Interval]
-----+-----
              |
    age_decades |   .185762   .0948398     1.96   0.057  -.005375   .376899
    _cons       |  -.2202967   .5902497   -0.37   0.711  -1.409867   .9692734
-----+-----
```

APPENDIX II: Diagnostic plot for Problem 1



APPENDIX III: Scatterplot for Problem 2



APPENDIX IV: Stata output for Problem 6

```
. regress sbp alcoh i.race, robust
```

```
Linear regression                Number of obs   =       735
                                F(4, 730)      =       2.86
                                Prob > F            =       0.0228
                                R-squared            =       0.0162
                                Root MSE          =       19.556
```

	sbp	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
	alcoh	.365078	.1571445	2.32	0.020	.0565689	.673587
	race						
	2	4.508028	2.175304	2.07	0.039	.2374294	8.778626
	3	-.2145398	2.962762	-0.07	0.942	-6.03109	5.60201
	4	6.091013	6.691777	0.91	0.363	-7.046412	19.22844
	_cons	129.6117	.8431304	153.73	0.000	127.9564	131.2669

```
. testparm i.race
```

- (1) 2.race = 0
- (2) 3.race = 0
- (3) 4.race = 0

```
F( 3, 730) = 1.68
Prob > F = 0.1693
```


APPENDIX V: Stata output for Problem 7

```
. regress sbp c.alcoh##i.race, robust
```

```
Linear regression                Number of obs   =       735
                                F(7, 727)       =         1.70
                                Prob > F             =       0.1049
                                R-squared            =       0.0175
                                Root MSE         =       19.584
```

	sbp	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
	alcoh	.4325421	.1891784	2.29	0.023	.061141	.8039431
	race						
	2	4.977631	2.418321	2.06	0.040	.2299052	9.725356
	3	.2108627	3.197077	0.07	0.947	-6.065743	6.487469
	4	8.523285	7.436498	1.15	0.252	-6.076288	23.12286
	race#c.alcoh						
	2	-.2112467	.425641	-0.50	0.620	-1.046879	.6243856
	3	-.2397076	.3406035	-0.70	0.482	-.9083915	.4289763
	4	-.7565144	.86767	-0.87	0.384	-2.459952	.9469235
	_cons	129.4707	.8654675	149.60	0.000	127.7716	131.1698

```
. testparm c.alcoh race#c.alcoh
```

- (1) alcoh = 0
- (2) 2.race#c.alcoh = 0
- (3) 3.race#c.alcoh = 0
- (4) 4.race#c.alcoh = 0

```
F( 4, 727) = 1.54
Prob > F = 0.1877
```