

# **The Central Limit Theorem**

## The Bridge Between Probability and Statistics

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# Outline

## **The Central Limit Theorem**

- Normal approximation to the binomial distribution
- Sampling distributions

# Outline

## Beware!

- If you've taken a statistics course, you *may* have heard of a law that says something about things becoming approximately normally distributed as sample sizes get large.
- **Forget what you think you've heard about that rule!**
  - We are about to learn it the *right* way!

# Review of Bernoulli distribution

## Recall the Bernoulli distribution (binary variables)

- The prevalence of some characteristic or trait (e.g., BRCA mutation) in the population is  $p$ .
- Randomly sample a single individual from that population; let  $X$  be 1 if that person has this trait, and 0 otherwise.
- Then,  $X \sim \text{Bernoulli}(p)$ , so that  $X$  takes on the value 1 with probability  $p$  and the value 0 with probability  $1 - p$ .

# Review of binomial distribution

## Recall

- The prevalence of some characteristic or trait (e.g., BRCA mutation) in the population is  $p$ .
- Now, sample  $n$  people and record 0/1 for each depending on whether they have the trait. Each of those  $n$  random variables,  $X_1, \dots, X_n$ , has a Bernoulli( $p$ ) distribution.
- So,  $T = \sum_{i=1}^n X_i = X_1 + \dots + X_n \sim \text{Binomial}(n, p)$ .
- **Two key points:**
  - ①  $T$  is the **sum** of  $n$  *independent and identically distributed* (iid) random variables.
  - ② Obtaining a value for  $T$  can be thought of as conducting a single study, *or* conducting  $n$  independent “mini-studies.”

# Review of binomial distribution

## Limiting behavior

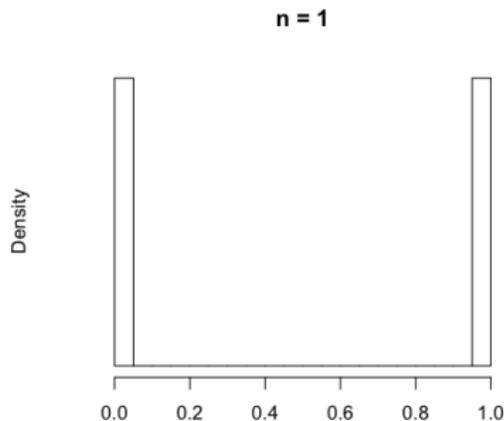
- So, suppose that  $X_1, \dots, X_n \sim \text{Bernoulli}(p)$  represent  $n$  *independent and identically distributed* random variables (e.g., the result of sampling people from a population with mutation prevalence  $p$ ).
- $T = \sum_{i=1}^n X_i = X_1 + \dots + X_n \sim \text{Binomial}(n, p)$ .
  - Randomly sample  $n$  individuals from the target population and *count* the number of those  $n$  who have the trait/characteristic; that count has a binomial distribution.
- It turns out that if  $n$  is very large, then  $T$  has an approximate normal distribution.
  - Let's take a look!

## Example: Prostate cancer cells

### Binomial distribution!

- The Binomial( $n, p$ ) distribution has two parameters ( $n$  and  $p$ )!
- Let us imagine that 50% of men over the age of 60 would test positive for the presence of prostate cancer cells ( $p = 0.5$ ).
- We want to see what happens when we sample  $n$  people from this population and count the number of men who test positive for the presence of prostate cancer cells.

## Binomial distribution: Sample one individual

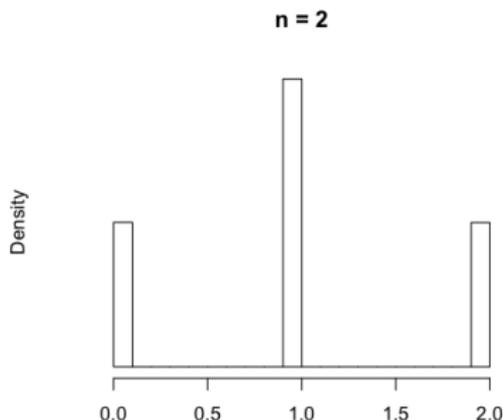


Binomial( $n = 1, p = 0.5$ )

*(Half the time, you would sample an individual with prostate cancer cells, and half the time, you would sample an individual without.)*

## Binomial distribution: Sample two individuals

How many have prostate cancer cells?

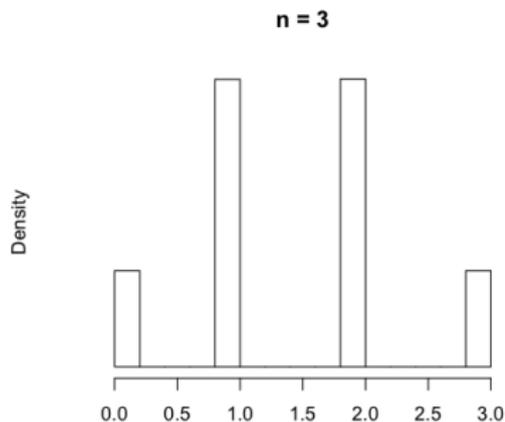


Binomial( $n = 2, p = 0.5$ )

*(25% of the time, neither would; 50% of the time, exactly one would; and 25% of the time, both would.)*

## Binomial distribution: Sample three individuals

How many have prostate cancer cells?

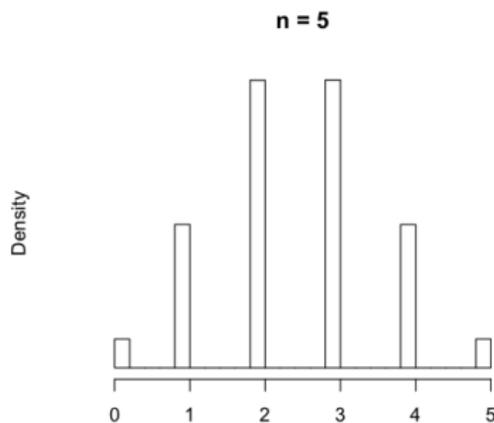


Binomial( $n = 3, p = 0.5$ )

*(12.5% of the time, none; 37.5% of the time, exactly one; 37.5% of the time, exactly two; and 12.5% of the time, all three.)*

## Binomial distribution: Sample five individuals

How many have prostate cancer cells?

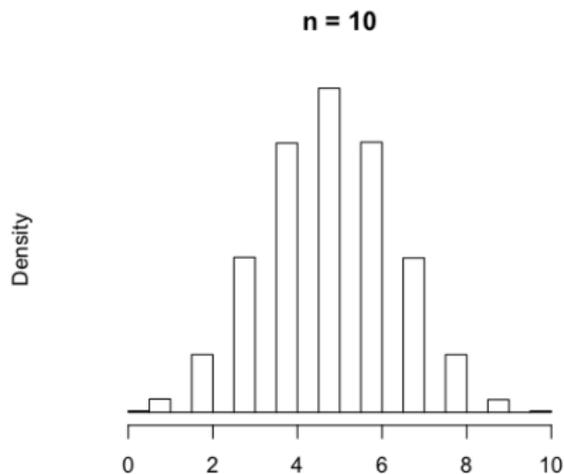


Binomial( $n = 5, p = 0.5$ )

(... and so on...)

# Binomial distribution: Sample ten individuals

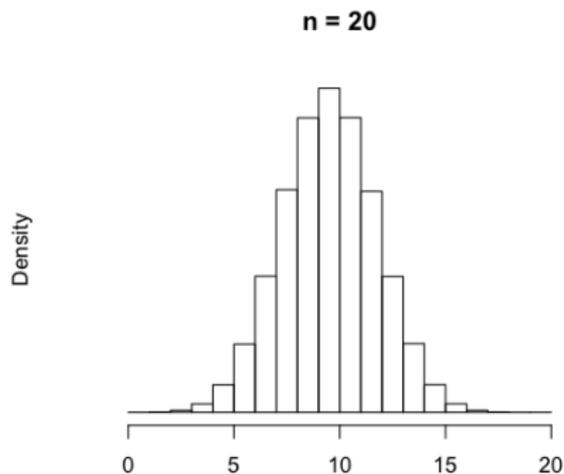
How many have prostate cancer cells?



Binomial( $n = 10, p = 0.5$ )

# Binomial distribution: Sample twenty individuals

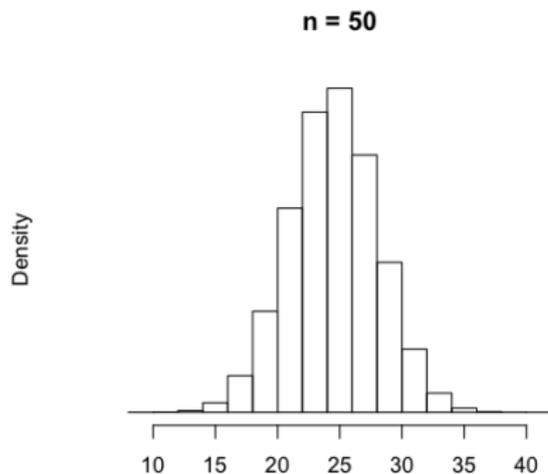
How many have prostate cancer cells?



Binomial( $n = 20, p = 0.5$ )

# Binomial distribution: Sample fifty individuals

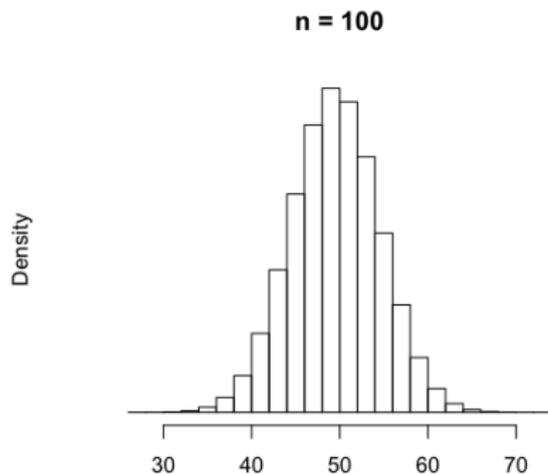
How many have prostate cancer cells?



Binomial( $n = 50, p = 0.5$ )

# Binomial distribution: Sample one-hundred individuals

How many have prostate cancer cells?



Binomial( $n = 100, p = 0.5$ )

# Binomial distribution

## Thoughts?

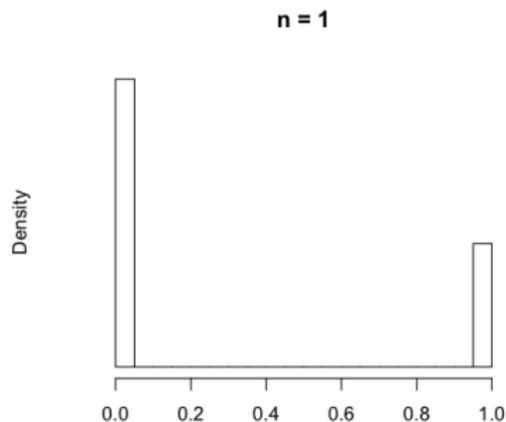
- In this example,  $\text{Binomial}(n, p = 0.5)$  is always symmetric, regardless of what  $n$  is.
- I think we can all agree that as  $n$  gets larger, appears more like a normal distribution.
- But *which* normal distribution? Suppose  $n = 100$ :
  - In truth,  $E[T] = np = 50$  and  $\text{Var}[T] = np(1 - p) = 25$ : one *hopes* that if  $T$  “looks normal,” that it would look like the normal distribution of mean 50 and variance 25.
  - It does!
- Let's make sure Andrew isn't just making things up and try another example.

## Example: Hypertension

### Binomial distribution!

- The Binomial( $n, p$ ) distribution has two parameters ( $n$  and  $p$ )!
- Let us imagine that 30% of people over the age of 50 suffer from hypertension ( $p = 0.3$ ).
- We want to see what happens when we sample  $n$  people from this population and count the number of people who have hypertension.

## Binomial distribution: Sample one individual

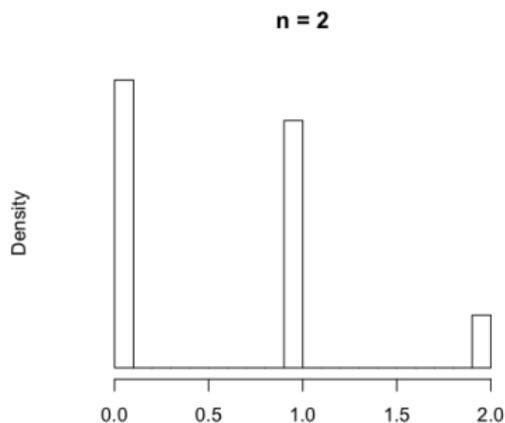


Binomial( $n = 1, p = 0.3$ )

*(30% of the time, would sample individual with hypertension; 70% of the time, would sample individual without.)*

## Binomial distribution: Sample two individuals

How many have hypertension?

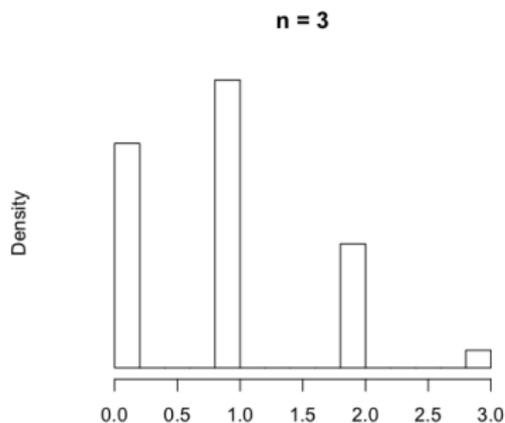


$\text{Binomial}(n = 2, p = 0.3)$

*(49% of the time, neither would; 42% of the time, exactly one would; and 9% of the time, both would.)*

## Binomial distribution: Sample three individuals

How many have hypertension?

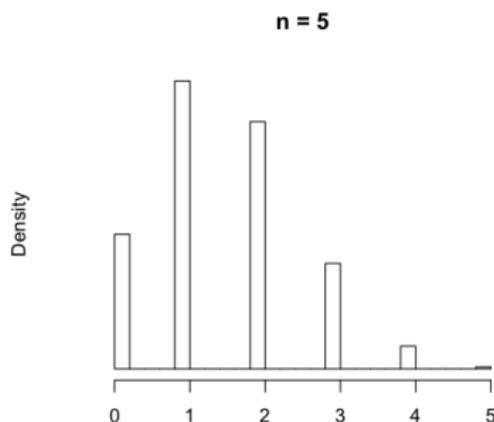


Binomial( $n = 3, p = 0.3$ )

*(34.3% of the time, none; 44.1% of the time, exactly one; 18.9% of the time, exactly two; and 2.7% of the time, all three.)*

## Binomial distribution: Sample five individuals

How many have hypertension?



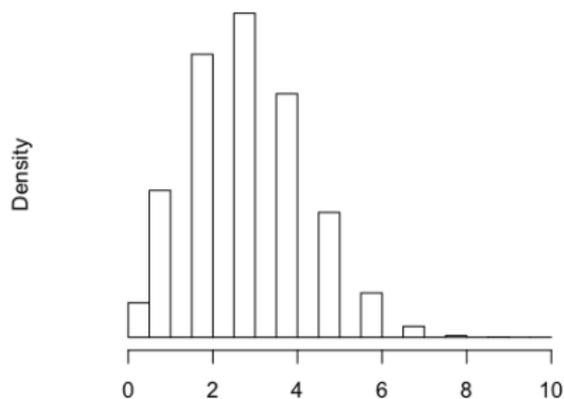
Binomial( $n = 5, p = 0.3$ )

(... and so on...)

# Binomial distribution: Sample ten individuals

How many have hypertension?

$n = 10$

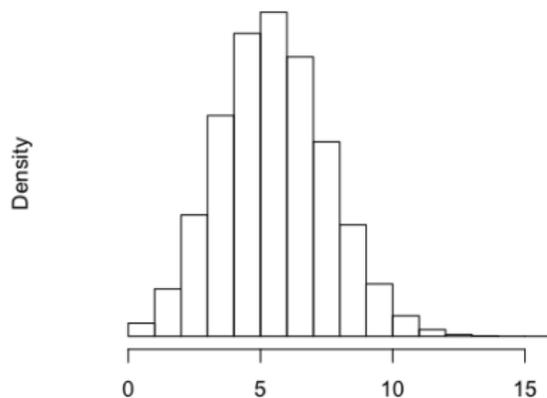


Binomial( $n = 10, p = 0.3$ )

# Binomial distribution: Sample twenty individuals

How many have hypertension?

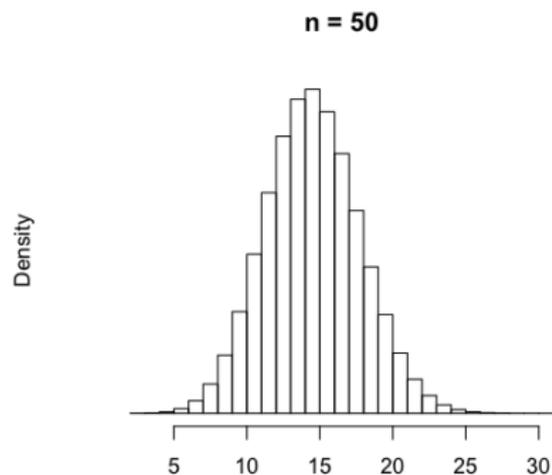
$n = 20$



Binomial( $n = 20, p = 0.3$ )

# Binomial distribution: Sample fifty individuals

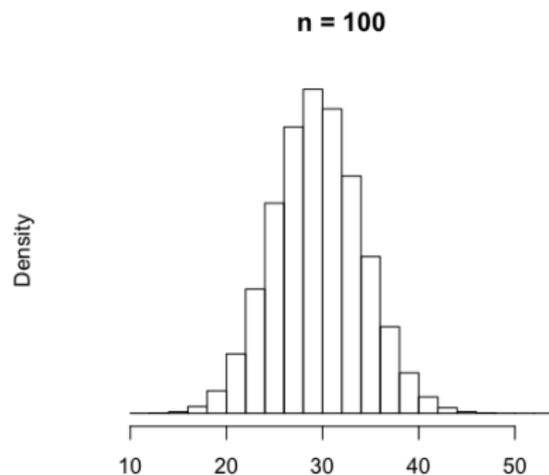
How many have hypertension?



Binomial( $n = 50, p = 0.3$ )

# Binomial distribution: Sample one-hundred individuals

How many have hypertension?



Binomial( $n = 100, p = 0.3$ )

# Binomial distribution

## Thoughts?

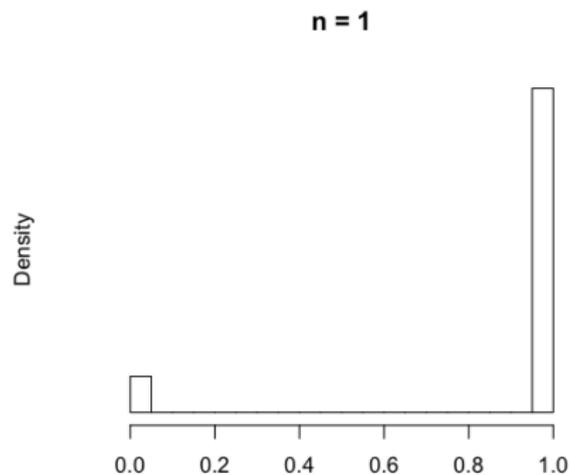
- In this example,  $\text{Binomial}(n, p = 0.3)$  is not generally symmetric; generally, appears to be a bit right-skewed with lower sample sizes.
- As  $n$  gets larger, the distribution becomes more symmetric. In fact, it starts to look like a normal distribution!
- But *which* normal distribution? Suppose  $n = 100$ :
  - In truth,  $E[T] = np = 30$  and  $\text{Var}[T] = np(1 - p) = 21$ : one *hopes* that if  $T$  “looks normal,” that it would look like the normal distribution of mean 30 and variance 21.
  - It does!
- Just one more time? Let's go wild and try  $p = 0.9$ .

## Example: Smoking and lung cancer

### Binomial distribution!

- The Binomial( $n, p$ ) distribution has two parameters ( $n$  and  $p$ )!
- Let us imagine that 90% of people under the age of 60 with small-cell lung cancer are smokers ( $p = 0.9$ ).
- We want to see what happens when we sample  $n$  people from this population and count the number of smokers.

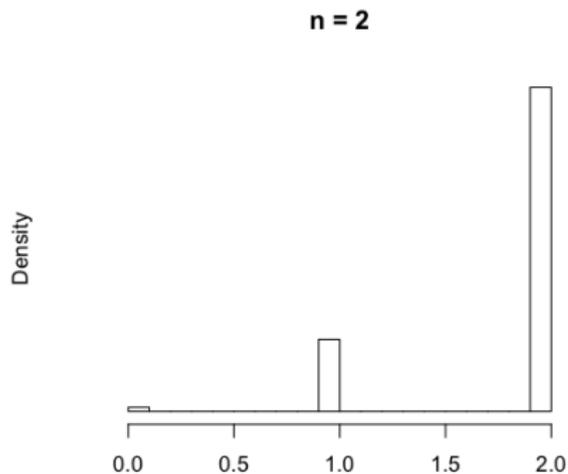
## Binomial distribution: Sample one individual



Binomial( $n = 1, p = 0.9$ )

*(90% of the time, would sample a smoker; 10% of the time, would sample non-smoker.)*

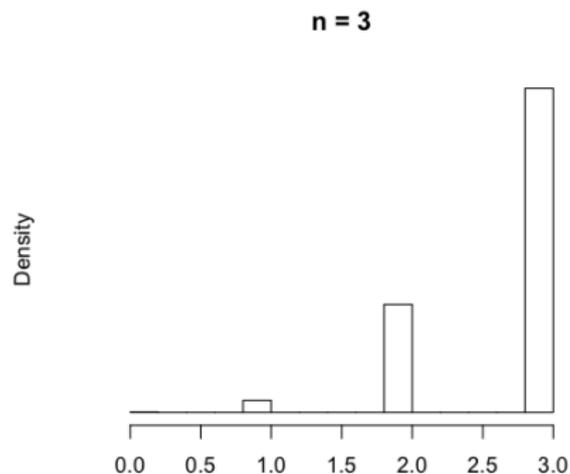
## Binomial distribution: Sample two individuals



Binomial( $n = 2, p = 0.9$ )

*(1% of the time, neither would; 18% of the time, exactly one would; and 81% of the time, both would.)*

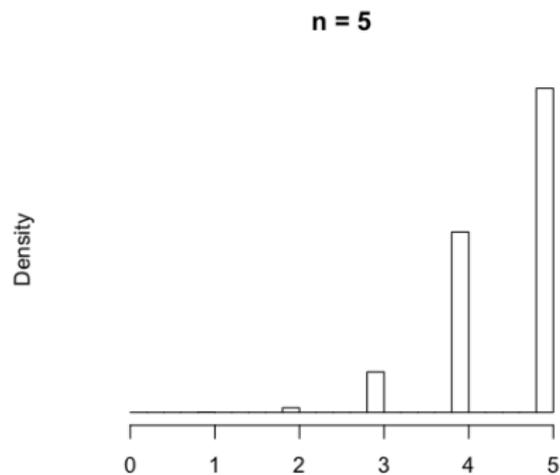
## Binomial distribution: Sample three individuals



Binomial( $n = 3, p = 0.9$ )

*(0.1% of the time, none; 2.7% of the time, exactly one; 24.3% of the time, exactly two; and 72.9% of the time, all three.)*

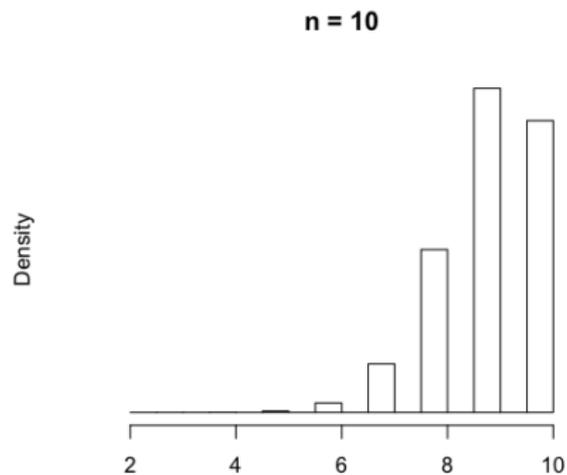
## Binomial distribution: Sample five individuals



Binomial( $n = 5, p = 0.9$ )

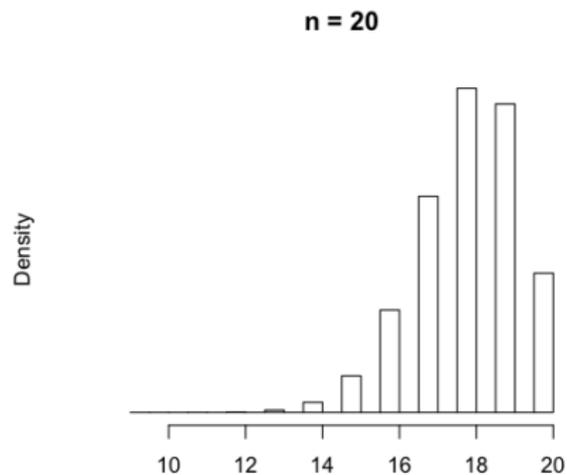
(... and so on...)

## Binomial distribution: Sample ten individuals



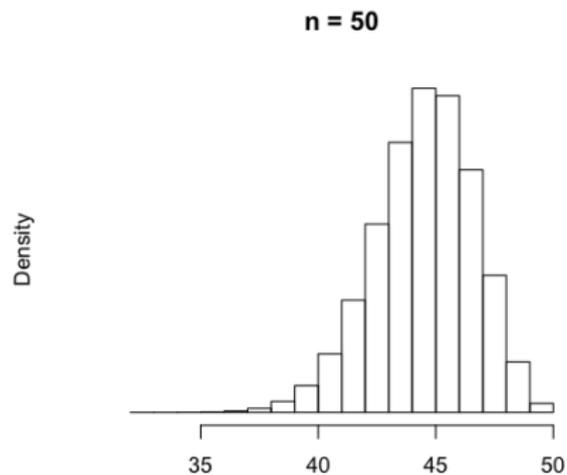
Binomial( $n = 10, p = 0.9$ )

# Binomial distribution: Sample twenty individuals



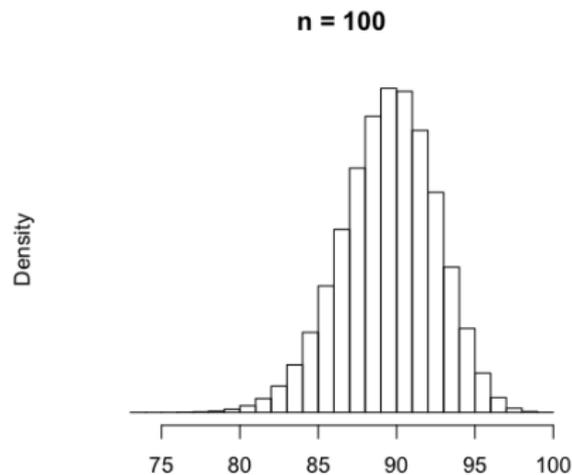
Binomial( $n = 20, p = 0.9$ )

# Binomial distribution: Sample fifty individuals



Binomial( $n = 50, p = 0.9$ )

# Binomial distribution: Sample one-hundred individuals



Binomial( $n = 100, p = 0.9$ )

# Binomial distribution

## Thoughts?

- In this example,  $\text{Binomial}(n, p = 0.9)$  is not generally symmetric; generally, appears to quite left-skewed with lower sample sizes.
- As  $n$  gets larger, the distribution becomes more symmetric. In fact, it starts to look like a normal distribution!
- But *which* normal distribution? Suppose  $n = 100$ :
  - In truth,  $E[T] = np = 90$  and  $\text{Var}[T] = np(1 - p) = 9$ : one *hopes* that if  $T$  “looks normal,” that it would look like the normal distribution of mean 90 and variance 9.
  - It does!

# The statement

## What's going on?

- It turns out that for large sample sizes, the Binomial distribution can be approximated by a normal distribution.
- Specifically, if  $X \sim \text{Binomial}(n, p)$ , if  $n$  is large enough, then:

$$X \dot{\sim} \mathcal{N}(np, np(1-p))$$

- Recall: We use symbol “ $\dot{\sim}$ ” to mean “is approximately distributed as,” as opposed to the symbol “ $\sim$ ”, which means “is exactly distributed as.”

# Normal approximation to the binomial distribution

## Example: Stage II bladder cancer

- For Stage II bladder cancer, the 5-year relative survival rate is approximately 63%
- You randomly sample 250 individuals with Stage-II bladder cancer and, at the end of five years, determine the number who are still alive. Let  $X$  denote the number still alive.
- Therefore,  $X \sim \text{Binomial}(n = 250, p = 0.63)$ .
- Why would we not want to compute  $P(X > 150)$  “by hand”?
  - Because, we don't want to have to evaluate the probability mass function 151 times.
- Exercise: Use the normal approximation to the binomial distribution to approximate  $P(X > 150)$ .

# Normal approximation to the binomial distribution

## Example: Stage II bladder cancer

- Here,  $X \sim \text{Binomial}(n = 250, p = 0.63)$ .
- We want to compute  $P(X > 150)$ .
  - $X \dot{\sim} \mathcal{N}(\mu = 250 \times 0.63, \sigma^2 = 250 \times 0.63 \times 0.37)$ .
  - $X \dot{\sim} \mathcal{N}(\mu = 157.5, \sigma^2 = 58.275)$ .
  - Recall:  $Z = (X - 157.5)/\sqrt{58.275} \dot{\sim} \mathcal{N}(0, 1)$ .
  - $P(Z > (150 - 157.5)/\sqrt{58.275}) = P(Z > -0.928) = 0.837$ .
  - NB: The true answer, using statistical software, is 0.821, not very far off from our approximation. The higher the sample size, the better the approximation.

# The reason

## Key point

- Binomial dist. can be approximated by normal dist. of the same mean and variance if sample size is large.
- Phenomenon doesn't occur for just *any* random variable!
- So, why does this happen in *this* case?
  - Because a binomial random variable can be expressed as a *sum* of independent, identically distributed (iid) random variables.
- Specifically,  $T = \sum_{i=1}^n X_i$ , where  $X_i \sim \text{Bernoulli}(p)$ .
- There is a statistical “rule” that says that sums of *iid* random variables will tend to have an approximate normal distribution the sample size is large (central limit theorem).

# Normal approximations to sums

## The central limit theorem

- Suppose  $X_1, \dots, X_n$  are iid with mean  $\mu$  and variance  $\sigma^2$ .
- Then, if  $n$  is “large enough,” then:

$$T = \sum_{i=1}^n X_i \sim \mathcal{N}(n\mu, n\sigma^2)$$

- This is true *even* when the individual  $X$ 's are not themselves sampled from a normal distribution! That's the magic of the theorem.

## Other facts

### Normal approximation to negative binomial distribution

- If  $X \sim \text{NegBinomial}(k, p)$ , and if  $k$  is large enough:

$$X \dot{\sim} \mathcal{N}\left(\mu = \frac{k}{p}, \sigma^2 = \frac{k(1-p)}{p^2}\right)$$

- Why does this happen? Because  $X$  can be expressed as the sum of  $k$  *iid*  $\text{Geometric}(p)$  random variables!

## Other facts

### Normal approximation to Poisson distribution

- If  $X \sim \text{Poisson}(\lambda)$ , and if  $\lambda$  is large enough:

$$X \dot{\sim} \mathcal{N}(\mu = \lambda, \sigma^2 = \lambda)$$

- Why does this happen? Because  $X$  can be expressed as the sum of  $n$  iid  $\text{Poisson}(\lambda/n)$  random variables!

**Normal approximations to distributions other than the binomial distribution often appear as optional problems on homework. 😊**

# How else is this useful?

## Using the central limit theorem

- Sample  $n = 100$  people and record their LDL values.
- These LDL values are random variables: each takes on a single value as result of a random sampling process.
- If  $X_i$  denotes the LDL value for subject  $i$ , then the *sample mean*,  $\bar{X}$  is, too, a random variable:

$$\bar{X} = \frac{1}{100} \sum_{i=1}^{100} X_i$$

- What does it mean for  $\bar{X}$  to be a random variable?
- In a single study,  $\bar{X}$  takes on a value,  $\bar{x}$  (of many possible values), as the result of a random sampling process.
  - Should we do this study again, would get a different value for  $\bar{X}$ . And again, would get something different from other two.

## How else is this useful?

### Using the central limit theorem

- If  $X_i$  denotes the LDL value for subject  $i$ , then the *sample mean*,  $\bar{X}$  is, too, a random variable:

$$\bar{X} = \frac{1}{100} \sum_{i=1}^{100} X_i$$

- Point: If central limit theorem tells you that sums of *iid* random variables are approximately normally distributed, then the sample mean is approximately normally distributed.
- Why? Because if,  $T = \sum_{i=1}^n X_i \dot{\sim} \mathcal{N}(n\mu, n\sigma^2)$ , then:

$$\bar{X} = T/n \dot{\sim} \mathcal{N}(\mu, \sigma^2/n).$$

- **I am equally interested in your ability to interpret what this means as I am in your ability to apply this!**

# Sampling distribution of the mean

## Recall:

- $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$ , the sample mean, is a random variable.
  - The sample mean LDL from a study of  $n = 100$  people, for example. It takes on one of many, many possible values in a given study.
- $\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$  is the sample mean, a *statistic*, computed from one specific data set.
  - $\bar{x}$  denotes the specific value of the sample mean computed from your single study of, for example,  $n = 100$  people.  $\bar{x} = 120$   $\mu\text{g}/\text{dL}$ , for instance.

# Sampling distribution of the mean

## For clarity:

- $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$  is a random variable with a distribution.
- $\bar{x}$  is computed from a single study, and takes on one value of *many* possible values it could have taken on.

Study number ( $k$ )	$\bar{X}$
$k = 1$	$\bar{x}_1$
$k = 2$	$\bar{x}_2$
$k = 3$	$\bar{x}_3$
$\vdots$	$\vdots$

# Sampling distribution of the mean

## Applying the central limit theorem

- Let  $X_1, \dots, X_n$  the LDL values for  $n$  randomly sampled individuals (assume  $n$  is large).
- If I were to conduct the above study in the same way, repeatedly, each time recording the sample mean, what would the distribution of those sample means look like?
- **Answer:**

$$\bar{X} \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right).$$

- The beauty here is that we don't even need to know the distribution of LDL! For large enough samples, we know (approximately) the distribution of the sample means.

# Mean and variance of sample mean

## Recall:

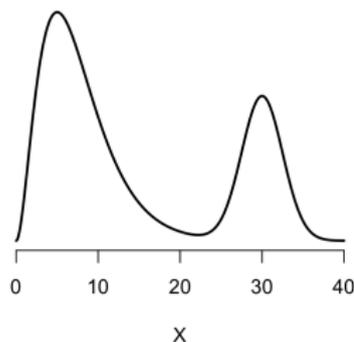
- Some math in an earlier set of lecture notes showed us that:
  - ① The sample mean,  $\bar{X}$ , is *unbiased* for the population mean,  $\mu$ .
  - ② The sample mean,  $\bar{X}$  gets closer and closer to  $\mu$  the larger your sample size,  $n$ , becomes larger.
- You're not expected to remember or replicate the math (it's there for your reference), but the two concepts above are important to understand.
- **Discussion point:** How does the central limit theorem square with the two points above?

# Sampling distribution of the mean

## Example: Something more wacky?

- Sampling a wacky continuous distribution—say, net insurance claims (in thousands of dollars).

A Wacky Distribution!



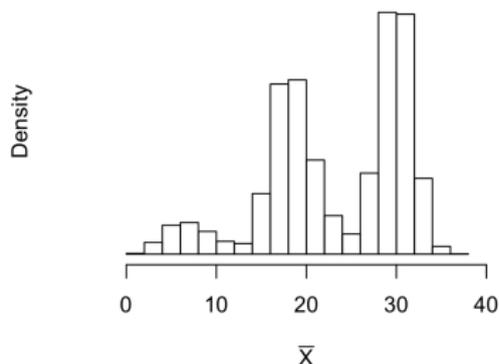
- NB: If  $n = 1$ , then the sampling distribution of  $\bar{X}$  is the distribution of  $X$  (make sure this makes sense)!

# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, X_2 \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means ( $n = 2$ )

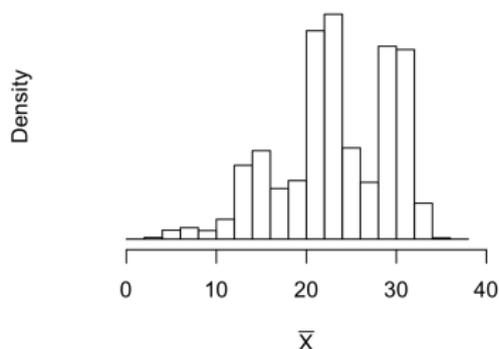


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, X_2, X_3 \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means ( $n = 3$ )

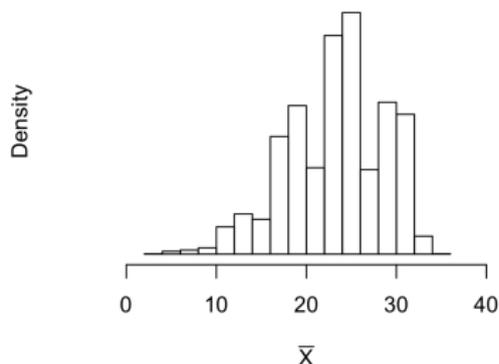


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_4 \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means ( $n = 4$ )

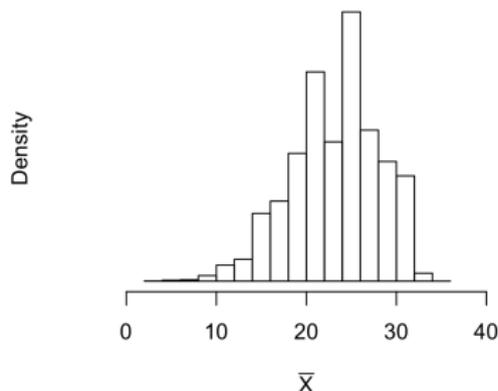


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_5 \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means ( $n = 5$ )

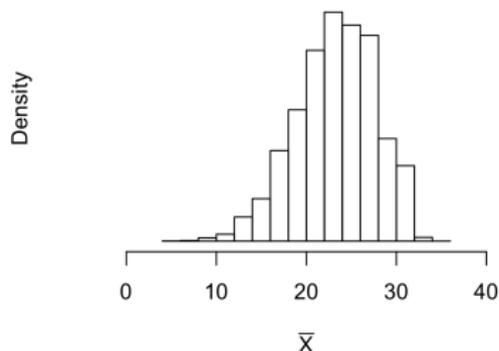


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_6 \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means (n = 6)

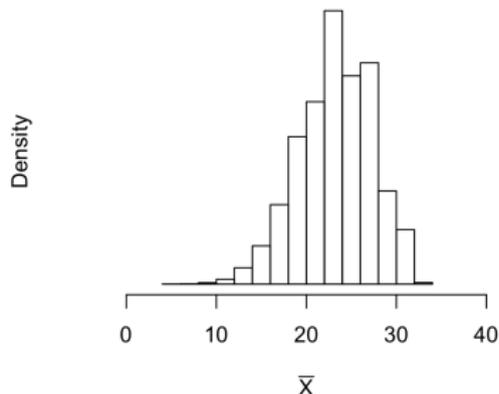


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_7 \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means ( $n = 7$ )

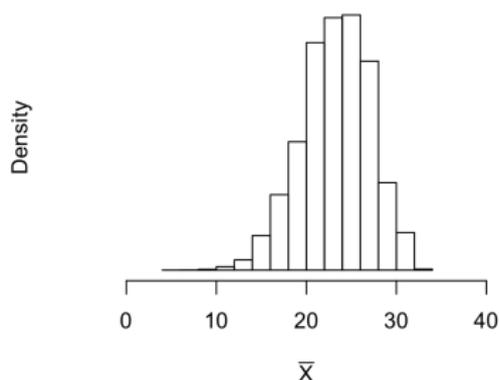


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_8 \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means ( $n = 8$ )

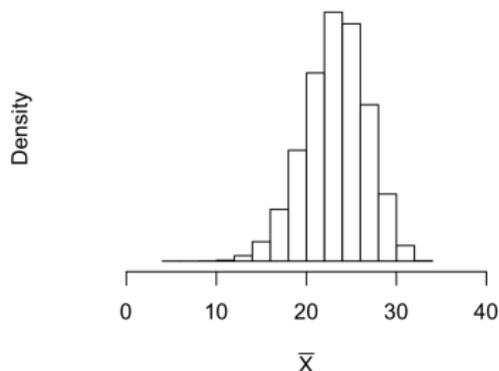


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_{10} \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means ( $n = 10$ )

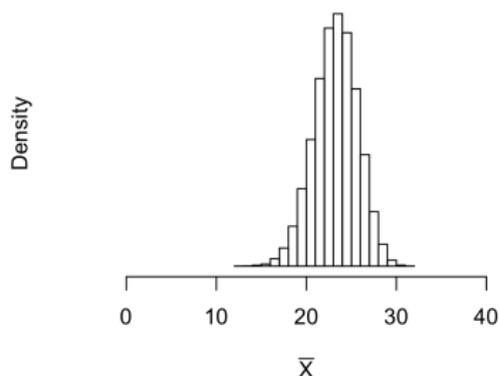


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_{20} \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means (n = 20)

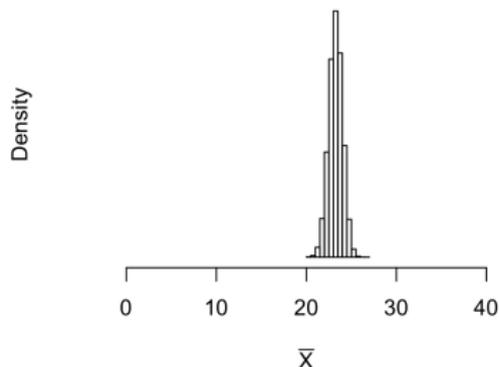


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_{200} \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means (n = 200)

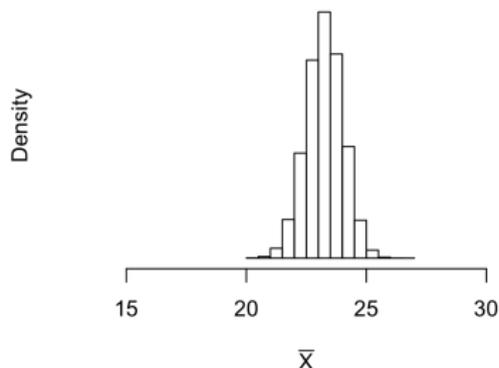


# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_{200} \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means (n = 200)



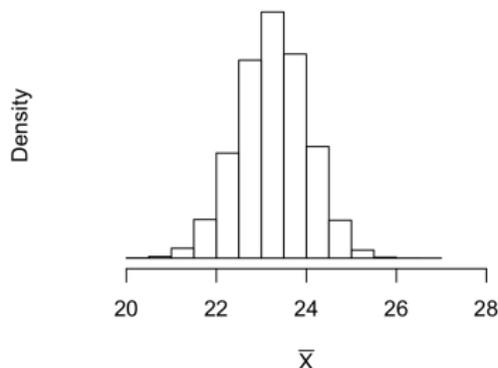
- Same figure, but reducing the x-axis range.

# Sampling distribution of the mean

## Example: Something more wacky?

- Example: Suppose  $X_1, \dots, X_{200} \sim$  Wacky Distribution. What is the *sampling* distribution of  $\bar{X}$ ?

Histogram of Sample Means (n = 200)



- Same figure, but reducing the x-axis range further.

# Sampling distribution of the mean

## Believe me yet? 😊

- It really does seem that, even if the distribution of your variable of interest is skewed and/or multimodal (or otherwise wacky), whether it is:
  - LDL
  - Insurance claims
  - Blood pressure
  - Cognitive abilities screening instrument
  - Number of hours in ICU
- ... there is **nothing** you can do to escape the following fact:
  - **If your sample size is large enough, the distribution of the sample means over study replicates will be approximately normally distributed.**
    - In particular, with mean  $\mu$  and variance  $\sigma^2/n$ .

# Sampling distribution of the mean

## Understanding the theorem:

- Sample  $X_1, \dots, X_n$  for  $n = 1$ . In this case,  $\bar{x} = x_1$ .
- If I repeated this study “infinitely” many times, what would the distribution of  $\bar{X}$  look like?

Study number ( $k$ )	$n = 1$
$k = 1$	$\bar{x}_1$
$k = 2$	$\bar{x}_2$
$k = 3$	$\bar{x}_3$
$\vdots$	$\vdots$
Dist. of $\bar{X}$	Same as dist. of $X$

# Sampling distribution of the mean

## Understanding the theorem:

- Sample  $X_1, \dots, X_n$  for  $n =$  “small” .
- If I repeated this study “infinitely” many times, what would the distribution of  $\bar{X}$  look like?

Study number ( $k$ )	$n =$ “small”
$k = 1$	$\bar{x}_1$
$k = 2$	$\bar{x}_2$
$k = 3$	$\bar{x}_3$
$\vdots$	$\vdots$
Dist. of $\bar{X}$	???

# Sampling distribution of the mean

## Understanding the theorem:

- Sample  $X_1, \dots, X_n$  for  $n =$  “medium”.
- If I repeated this study “infinitely” many times, what would the distribution of  $\bar{X}$  look like?

Study number ( $k$ )	$n =$ “medium”
$k = 1$	$\bar{X}_1$
$k = 2$	$\bar{X}_2$
$k = 3$	$\bar{X}_3$
$\vdots$	$\vdots$
Dist. of $\bar{X}$	Closer to normal than small sample

# Sampling distribution of the mean

## Understanding the theorem:

- Sample  $X_1, \dots, X_n$  for  $n = \text{"huge"}$ .
- If I repeated this study "infinitely" many times, what would the distribution of  $\bar{X}$  look like?

Study number ( $k$ )	$n = \text{"huge"}$
$k = 1$	$\bar{X}_1$
$k = 2$	$\bar{X}_2$
$k = 3$	$\bar{X}_3$
$\vdots$	$\vdots$
Dist. of $\bar{X}$	Approximately normal*

\*Theorem asserts that there is an  $n$  large enough that this will be the case.

# The central limit theorem

## Formally: The central limit theorem!

- Suppose  $X_1, \dots, X_n$  are independently sampled from a common distribution with mean  $\mu$  and variance  $\sigma^2$ . As  $n$  grows larger and larger,

$$P\left(\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} < s\right) \rightarrow P(Z < s) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^s e^{-\frac{1}{2}t^2} dt,$$

... where  $Z \sim \mathcal{N}(0, 1)$ .

- **You do not need to remember/use this formula.**

# The central limit theorem

## Ways of stating the central limit theorem!

- Suppose  $X_1, \dots, X_n$  are *iid* with common mean  $\mu$  and variance  $\sigma^2$ . Then, for large sample sizes,  $n$ :

$$\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \quad \dot{\sim} \quad \mathcal{N}(0, 1), \text{ or}$$

$$\sqrt{n}(\bar{X} - \mu) \quad \dot{\sim} \quad \mathcal{N}(0, \sigma^2), \text{ or}$$

$$\bar{X} \quad \dot{\sim} \quad \mathcal{N}\left(\mu, \frac{\sigma^2}{n}\right)$$

# The central limit theorem

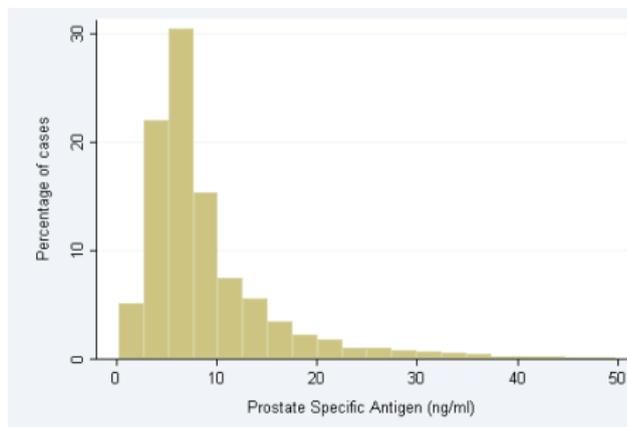
## Thoughts

- To me, this is wild for two reasons:
  - ① This works for any distribution, whether discrete or continuous. Whether symmetric or skewed. Whether the range is finite or infinite. Whether it's unimodal or bimodal.
  - ② How lucky we are that the sample mean looks more and more like a *normal* distribution—one that we understand so well.
- Caution: The central limit theorem does not say anything about the distribution of the variable itself as the sample size grows. The distribution of the variable you're measure *does not change* with sample size.

# The central limit theorem

## Application: PSA

- PSA (prostate specific antigen) a biomarker used to detect prostate cancer.
- Among men undergoing surgery for prostate cancer:
  - Mean: approximately 10 ng/mL.
  - Variance: approximately 11 ng/mL.



# The central limit theorem

## Application: PSA

- Among men undergoing surgery for prostate cancer:
  - Mean: approximately 10 ng/mL.
  - Variance: approximately 11 ng/mL.
- **Exercises:** Sample  $n = 120$  men undergoing surgery for prostate cancer, and record their PSA values.
  - You decide to plot a histogram of their PSA values. Do you believe that histogram would suggest that the distribution of the PSA values would be approximately normally distributed?
  - You compute the sample mean PSA value. With approx. what probability would it be greater than 11 ng/mL?
  - Approximately what values mark the 2.5th and 97.5th percentiles of the sampling distribution of  $\bar{X}$ ?

# The central limit theorem

## Application: PSA

- True mean PSA: approximately 10 ng/mL.
- True variance of PSA: approximately 11 ng/mL.
- Sample  $n = 120$  men undergoing surgery for prostate cancer, and record their PSA values.
- **Exercise:** You decide to plot a histogram of their PSA values. Do you believe that histogram would suggest that the distribution of the PSA values would be approximately normally distributed?
  - **Answer:** No!

# The central limit theorem

## Application: PSA

- True mean PSA: approximately 10 ng/mL.
- True variance of PSA: approximately 11 ng/mL.
- Sample  $n = 120$  men undergoing surgery for prostate cancer, and record their PSA values.
- **Exercise:** You compute the sample mean PSA value. With approx. what probability would it be greater than 10.3 ng/mL?
  - **Answer:** The central limit theorem asserts that

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim \mathcal{N}(0, 1).$$

- Here,  $z = (10.3 - 10)/(\sqrt{11/120}) = 0.9909$ .
- $P(Z > 0.9909) = 1 - 0.8391 = 0.161$ .

# The central limit theorem

## Application: PSA

- True mean PSA: approximately 10 ng/mL.
- True variance of PSA: approximately 11 ng/mL.
- Sample  $n = 120$  men undergoing surgery for prostate cancer, and record their PSA values.
- **Exercise:** Approximately what values mark the 2.5th and 97.5th percentiles of the sampling distribution of  $\bar{X}$ ?
  - **Answer:** The central limit theorem asserts that

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim \mathcal{N}(0, 1).$$

- Recall:  $\pm 1.96$  mark the 2.5th and 97.5th percentiles of standard normal dist. Can convert back to the PSA scale:
- $\mu \pm 1.96 \frac{\sigma}{\sqrt{n}} = 10 \pm 1.96 \times \sqrt{11/120} = [9.41 \text{ ng/mL}, 10.6 \text{ ng/mL}]$ .

# The central limit theorem

## Application: PSA

- Interpretation of activity results: If I were to conduct this study over, and over, and over, each time recording the sample mean, then:
  - 16.1% of those sample means would be greater than 10.3 ng/mL.
  - The 2.5th and 97.5th percentiles of the distribution of the sample means are approximately 9.41 ng/mL and 10.6 ng/mL.

# The central limit theorem

## Aside: Why the “normal” distribution?

- The normal distribution is maximally “disorderly”.
  - The *entropy* of a random variable  $X$  is  $E[\log X]$ .
  - If  $X$  is normally distributed with expectation  $\mu$  and variance  $\sigma^2$ , then *no other* random variable with expectation  $\mu$  and variance  $\sigma^2$  has higher entropy.
- Makes sense that if we randomly sample values from a distribution, the repeat-sample distribution of the sample mean should head toward the most *disorderly* distribution.

# The central limit theorem: The point?

## Keeping our eye on the prize!

- May wish to *estimate*, population mean  $\mu$ .
  - Know how to do this: compute  $\bar{x}$  in the sample.
- Want to quantify degree of precision with which we know  $\mu$ .
  - Rely on information about *sampling* distribution of  $\bar{X}$ : “what does the distribution of  $\bar{X}$  look like if I repeat this study?”
    - ① Typically, don't know *exact* sampling dist. of  $\bar{X}$ , and ...
    - ② Typically, cannot actually *perform* study over and over...
- **Central limit theorem (CLT) is our friend!** It is ultimately a statement about the approximate sampling distribution of  $\bar{X}$  for large sample sizes. You will use it for the rest of the year, and pretty much the rest of your research career.

# Sampling distribution of the mean

## Population vs. sample mean

- I know how to compute  $\bar{x}$  from a data set.
- With this information, I want to uncover some sort of information about  $\mu$ , the population mean.
- In our blood pressure example, this could mean asking, for example, the following questions:
  - What are the true values of the population average blood pressure with which my data are consistent?
    - **Confidence intervals! Characterize precision of estimate.**
  - If the true value of the population mean blood pressure were 130 mm Hg, with what frequency would I observe a sample mean at least as high as the one I observed?
    - **p-values! Strength of evidence in support of your hypothesis.**

# Sampling distribution of the mean

## Population vs. sample mean

- Ability to obtain answers to these sorts of questions regarding the population mean is inherently tied to our understanding of the *sampling distribution* of  $\bar{X}$ .
- That is, we need to ask ourselves the following question:
  - “If I were to complete this study over and over again, and each time compute  $\bar{x}$ , what would the distribution of the values of  $\bar{x}$  look like?”
- Reminder: We typically cannot name the distribution of  $\bar{X}$ .
  - If we *knew* the exact form of the distribution of  $\bar{X}$ , then in some sense this wouldn't be too hard of a problem.

# The central limit theorem

## Summary

- Major theorem!
- Gives us approximate sampling distribution of  $\bar{X}$  when the population parameters of  $X$  are known.
  - If I repeated the study over and over, recording the value  $\bar{x}$  each time, what would the distribution of those values be?  
With large samples, approximately normal!
- Building foundation to handle the case where the population parameters are *unknown*.
- In turn, we will shortly be able to answer questions about:
  - The precision of our estimate of the population mean.
  - The strength of our evidence for a hypothesis regarding the population mean.

I leave you with spooky statistics!

